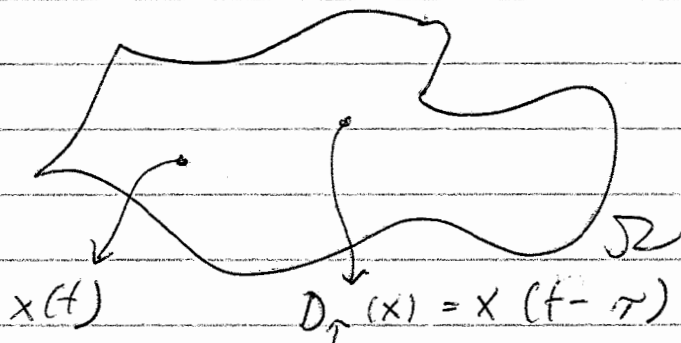


1. Power Spectra

The general problem: characterize a "noise" $x(t)$.



We typically have access to multiple samples at different start times. So, to get off the ground, assume

a) Stationarity: If $D_\tau(x) = x(t - \tau)$, then $p(D_\tau(x)) = p(x)$

b) Ergodicity $\langle \rangle_\Omega = \langle \rangle_\tau$

for "all" functions on $p(x), p(y), \dots$

c) For suff. large T , $x(t)$ and $x(t+T)$ are independent.

Note that defining $p(x)$ takes some care.

In discrete case, define

prob. ($x(t_1)$ betw $x_1 + x_1 + \delta x$ and

$x(t_2)$ " $x_2 + x_2 + \delta x$, "

\vdots

$x(t_k)$ " $x_k + x_k + \delta x$) = $P(x_1, \dots, x_k) (\delta x)^k$

2

and these prob's need to be defined in a self-consistent way, for all $\{x_1, \dots, x_k\}$, $k=1, 2, \dots$

Similarly in continuous time.

We can always take $\langle x(t) \rangle_{\mathcal{R}} = 0$, since

ergodicity \Rightarrow it is constant, \rightarrow can re-parametrize it $\langle x(t) \rangle_{\mathcal{R}} = \mu \neq 0$.

But of course sample mean may not be 0.

In view of results of last § (maxent + C.L.T.), we will take \mathcal{R} to be a Gaussian, but parameters of the Gaussian may not be known. Then, extend.

Since $\langle x(t) \rangle_{\mathcal{R}} = 0$, it is not a useful description

what about $C_x(\tau_1, \tau_2) = \langle x(\tau_1)x(\tau_2) \rangle_{\mathcal{R}}$?

Not $C_x(\tau_1, \tau_2) = \langle x(\tau_1+t)x(\tau_2+t) \rangle_{\mathcal{R}}$ (only t)

so $C_x(\tau_1+t, \tau_2+t) = C_x(\tau_1, \tau_2)$, and we can write

$$C_x(\tau_1, \tau_2) = C_x(\tau_1 - \tau_2), \text{ where } C_x(t) \equiv C_x(t, 0).$$

Also $C_x(\tau_1, \tau_2) = C_x(\tau_2, \tau_1)$, so $C_x(t) = C_x(-t)$.

This is the autocovariance. Autocorrel = $C_x(t)/C_x(0)$.

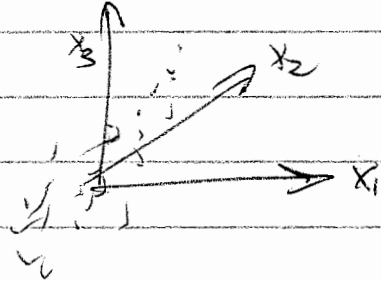
3.

Geometrically:

with $x = (x_1, x_2, \dots, x_n, \dots)$,

$C_x(t)$ describes covariance of x_r and x_{r+t} , i.e.,

the shape of the ellipse



But a) translation-in-time symmetry \Rightarrow certain covariances must be equal

b) Some covariances influence others, i.e.,

If x_3 and x_4 are highly correlated,
then x_4 and x_5 " " "
and x_3 and x_5 must be somewhat correlated

c) Some correlations appear easier to measure
(x_k, x_{k+1}) than others ($x_k, x_{k+large}$).

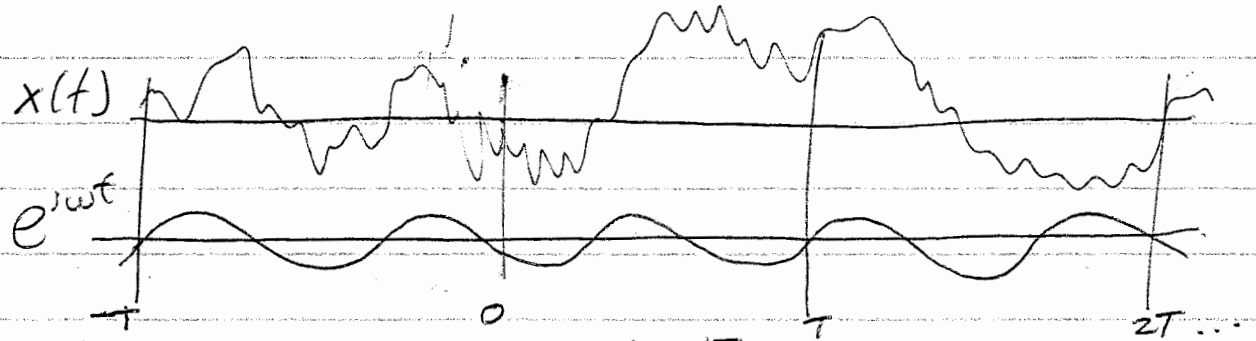
So it is not clear how to do estimations -
e.g., how to put error bars on $C_x(t)$

What about trying to characterize $\tilde{x}(\omega) = \int_{-\infty}^{\infty} x(t) e^{-j\omega t} dt$?

This should deal with time-translation. However;

$\tilde{x}(\omega)$ is ill-defined:

4.



$$I = \int_{-\infty}^{\infty} x(t) e^{-i\omega t} dt = \sum_{N=-\infty}^{\infty} \int_{NT}^{(N+1)T} x(t) e^{i\omega t} dt$$

$$= \sum_{N=-\infty}^{\infty} F(x; \omega, T, NT) = \sum_{N=-\infty}^{\infty} F_N$$

F_N = a "Fourier estimate" at freq. ω , length T ,
start time NT , of $x(t)$

How are the F_N 's distributed?

$$F(x; \omega, T, T_0) = \int_{T_0}^{T_0+T} x(t) e^{-i\omega t} dt$$

$$u = t + T_0$$

$$= \int_0^T x(u - T_0) e^{-i\omega(u - T_0)} du$$

$$= e^{i\omega T_0} \int_0^T x(u - T_0) e^{-i\omega u} du$$

$$= F(D_{T_0}(x); \omega, T, 0) \cdot e^{i\omega T_0}$$

5.

But \mathcal{D} is translation-invariant, so any statistic on x must = that statistic on $D_{T_0}(x)$.

So $F(x; \omega, T, T_0)$ must be distributed like

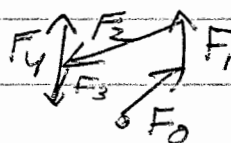
$$F(x; \omega, T, 0) \cdot e^{i\omega T_0}$$

But $F(x; \omega, T, T_0)$ must also be distributed like $F(x; \omega, T, 0)$ since when you set "wall clock" time to 0 is arbitrary.

So each $F(x; \omega, T, T_0)$ is distributed circularly-symmetrically in the complex plane.

But also, for T sufficiently large, "most" of the interval of $[NT, (N+1)T]$ is very far from the adjacent intervals, so "most" of each estimate F_N is independent of adjacent, & all other, intervals [condition (c)]

Take $I = \sum_{N=0}^{K-1} F_N$ but the F_N 's consist of random-walk steps in the plane:



So, $\langle I_{KT} \rangle = 0$ but

$$\langle |I_{KT}|^2 \rangle = \left\langle \left| \sum_{N=0}^{K-1} F_N \right|^2 \right\rangle = \sum_{N, M=0}^{K-1} \langle F_N \bar{F}_M \rangle = K \langle |F_0|^2 \rangle$$

6.

So $\frac{I}{KT} = \int_0^T x(t) e^{-i\omega t} dt$ has mean 0, and variance proportional to K (fixed T).
Proportionality depends on T .

Since T is arbitrary (but large),

I_{KT} must have variance proportional to KT .

Conversely, we expect that

$$\lim_{T \rightarrow \infty} \frac{1}{T} \left\langle \left| \int_0^T x(t) e^{-i\omega t} dt \right|^2 \right\rangle \text{ exists.}$$

This is the power spectrum. Units: $X^2 \cdot T$, or, X^2 / Hz

$$P_X(\omega) = \lim_{T \rightarrow \infty} \frac{1}{T} \left\langle \left| \int_0^T x(t) e^{-i\omega t} dt \right|^2 \right\rangle = \lim_{T \rightarrow \infty} \frac{1}{T} \left\langle |F(x; \omega, T, T_0)|^2 \right\rangle$$

Because of ergodicity (b), can replace $\langle \rangle_T$ by $\langle \rangle_{T_0}$.
Because of limited correlation length (c), can replace $\langle \rangle_{T_0}$ by

$$P_X(\omega) = \lim_{\substack{T \rightarrow \infty \\ N \rightarrow \infty}} \frac{1}{N} \sum_{n=0}^{N-1} \frac{1}{T} |F(x, \omega, T, nT + T_0)|^2$$

Note the difficulties with a "plus-in" estimator

- can't make $T \rightarrow \infty$
- with data of length L , reciprocity between length of segment T , & max # of segments L/T
- tradeoff betw frequency resolution $\Delta\omega$ and N , since $T \geq \frac{1}{\Delta\omega}$.

7.

What if there is a deterministic component in $x(t)$, namely,

$$x(t) = x_p(t) + x_N(t)$$

\swarrow "periodic" \nwarrow "noise"

$$F(x; \omega, T, T_0) = F(x_p, \omega, T, T_0) + F(x_N, \omega, T, T_0).$$

$F(x_N, \omega, T, T_0)$ behaves as above.

But $\langle F(x_p, \omega, T, T_0) \rangle$ typically not 0.

With $T = mP$, P a period of $x_p(t)$, and $\omega = \frac{2\pi h}{P}$

$$\begin{aligned} \langle F(x_p, \omega, T, T_0) \rangle &= \int_{T_0}^{mP+T_0} x_p(t) e^{-i \frac{2\pi h}{P} t} dt \\ &= m \int_{T_0}^{P+T_0} x_p(t) e^{-i \frac{2\pi h}{P} t} dt \\ &= T \cdot \frac{1}{P} \int_{T_0}^{P+T_0} x_p(t) e^{-i \frac{2\pi h}{P} t} dt \\ &= T (\tilde{x}_p)_h, \text{ i.e., proportional to } h^{-4n}. \\ &\quad \text{F.C. of } x_p. \end{aligned}$$

So, the defining limit of the P.S. does not exist.

8-

However, can write

$$P_{X_n}(\omega) = \lim_{T \rightarrow \infty} \frac{1}{T} \left\langle |F(x, \omega, T, T_0) - \langle F(x, \omega, T, T_0) \rangle|^2 \right\rangle_{\mathcal{R}}$$

Since $\langle F(x, \omega, T, T_0) \rangle$ typically not known,
it has to be estimated from data. Using standard
debiased variance estimate,

$$P_{X_n}(\omega) = \lim_{\substack{T \rightarrow \infty \\ N \rightarrow \infty}} \frac{1}{N-1} \sum_{n=0}^{N-1} \frac{1}{T} \left| F(x, \omega, T, nT+T_0) - \frac{1}{N} \sum_{n=0}^{N-1} F(x, \omega, T, nT+T_0) \right|^2$$

Why do we not need to consider the more general

$$W_X(\omega_1, \omega_2) = \lim_{T \rightarrow \infty} \frac{1}{T} \left\langle F(x; \omega_1, T, T_0) F(x; \omega_2, T, T_0) \right\rangle,$$

or even

$$W_X(\omega_1, \omega_2, \Delta T) = \lim_{T \rightarrow \infty} \frac{1}{T} \left\langle F(x; \omega_1, T, T_0) F(x; \omega_2, T, T_0 + \Delta T) \right\rangle?$$

Previous (p. 4-5) argument shows that shifting the
start-time T_0 by, say, ΔT and choosing a new
sample $D_{\Delta T}(x)$ multiplies the value of
 $F(x, \omega, T, T_0)$ by $e^{i\omega \Delta T}$.
Similarly, value of $F(x, \omega_2, T, T_0)$ is multiplied by $e^{i\omega_2 \Delta T}$.

So the product is multiplied by $e^{i(\omega_1 + \omega_2) \Delta T}$.

9.

So there will only be a nonzero expectation if, for all ωT ,

$$e^{i(\omega_1 + \omega_2)\Delta T} = 0, \text{ i.e., } \omega_1 + \omega_2 = 0, \text{ i.e.,}$$

$$W_X(\omega_1, -\omega_1) \equiv P_X(\omega) \quad \left[\lim_{T \rightarrow \infty} \frac{F(x, -\omega, T, T_0)}{F(x, \omega, T, T_0)} \right]$$

However, $B_X(\omega_1, \omega_2) \equiv$

$$\lim_{T \rightarrow \infty} \frac{1}{T} \langle F(x, \omega_1, T, T_0) \overline{F(x, \omega_2, T, T_0)} \overline{F(x, \omega_1 + \omega_2, T, T_0)} \rangle$$

make sense. "Bispectrum"

What about multichannel data?

Cross-spectrum:

$$C_{XY}(\omega) = \lim_{T \rightarrow \infty} \frac{1}{T} \langle F(x, \omega, T, T_0) \overline{F(y, \omega, T, T_0)} \rangle$$

No need to consider two frequencies.

If we can estimate P , we can estimate C_{XY} :

$$\begin{aligned} \langle (X+Y) \overline{(X+Y)} \rangle &= |X|^2 + |Y|^2 + X\bar{Y} + Y\bar{X} \\ &= |X|^2 + |Y|^2 + 2 \operatorname{Re}(X\bar{Y}) \end{aligned}$$

$$\begin{aligned} \langle (X+iY) \overline{(X+iY)} \rangle &= |X|^2 + |Y|^2 + iY\bar{X} - iX\bar{Y} \\ &= |X|^2 + |Y|^2 + 2 \operatorname{Im}(X\bar{Y}) \end{aligned}$$

10.

So,

$$\left. \begin{aligned} \operatorname{Re} C_{xy}(\omega) &= \frac{1}{2} [P_{x+y}(\omega) - P_x(\omega) - P_y(\omega)] \\ \operatorname{Im} C_{xy}(\omega) &= \frac{1}{2} [P_{x+iy}(\omega) - P_x(\omega) - P_y(\omega)] \end{aligned} \right\} \text{formal}$$

So an estimator for P will lead to an estimator for C ;
bias will be easy to calculate.

$$\text{Coherency} = \frac{C_{xy}(\omega)}{\sqrt{P_x(\omega)P_y(\omega)}} \quad \text{Complex \#, } | | \leq 1$$

$$\text{Coherence} = |\text{Coherency}|$$

Given $P_x(\omega)$, can one calculate a "typical" $x(t)$?

Draw from the max ent distrib specified by $P_x(\omega)$.

Naive approach: generate a signal of length L , starting at $t=0$,
sampled \hat{c} resolution Δt .

$F(x, \omega, 0, L)$ is a complex # $\hat{c} \langle |F|^2 \rangle = L P_x(\omega)$,
symmetrically distributed, i.e.,

$\operatorname{Re} F$ has variance $\frac{1}{2} L P_x(\omega)$

$\operatorname{Im} F$ " " $\frac{1}{2} L P_x(\omega)$

$\operatorname{Re} F$ + $\operatorname{Im} F$ are independently distributed

p.

So, choose $\text{Re } F(\omega)$ + $\text{Im } F(\omega)$ independently from

$$\frac{1}{\sqrt{2\pi V}} e^{-F^2/2V}, \quad V = \frac{1}{2} L P_x(\omega)$$

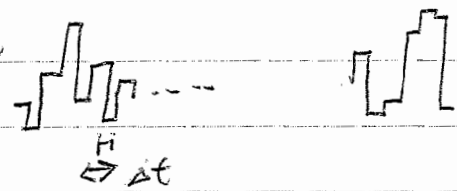
Do this for a sufficient set of ω so that one can write

$$x(t) = \frac{1}{\sqrt{L}} \sum_{\omega = \frac{1}{2\Delta t}}^{\frac{1}{2\Delta t}} F(\omega) e^{i\omega t}, \quad \text{constant value from } n\Delta t \text{ to } (n+1)\Delta t.$$

n is steps of $\frac{1}{L}$ $(N = \frac{L}{\Delta t})$

Note $F(\omega=0)$ or $F(\omega = \frac{1}{2\Delta t})$ must be real, &
 $F(-\omega) = \overline{F(\omega)}$.

The $\frac{1}{\sqrt{L}}$ is to ensure that $\int_0^L x(t) e^{-i\omega t} dt = \sqrt{L}$

This makes a signal that looks like 

- Problem 1: Its power spectrum, at intermediate freqs, doesn't agree w/ $P_x(\omega)$
2. It is periodic w/ period L .

Solution: use a larger L than you need. then $\Delta\omega$ is smaller.

12. Properties of the Power Spectrum

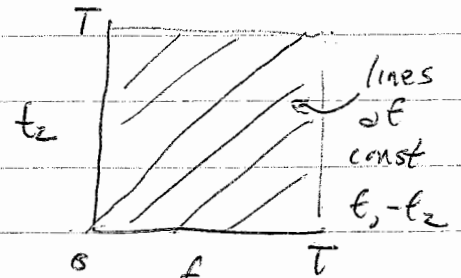
Relation to auto covariance

$$\text{From } P_x(\omega) = \lim_{T \rightarrow \infty} \frac{1}{T} \left\langle \left| \int_0^T x(t) e^{-j\omega t} dt \right|^2 \right\rangle$$

$$= \lim_{T \rightarrow \infty} \frac{1}{T} \left\langle \int_0^T \int_0^T x(t_1) x(t_2) e^{-j\omega t_1} e^{+j\omega t_2} dt_1 dt_2 \right\rangle$$

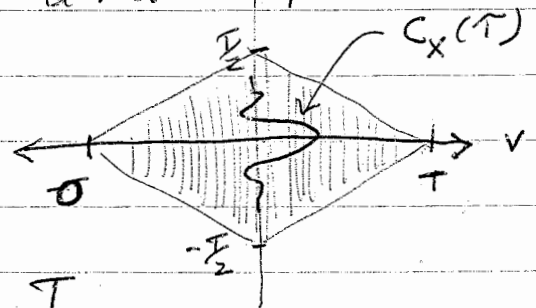
$$\tau = t_1 - t_2$$

$$v = \frac{1}{2}(t_1 + t_2)$$



$$P_x(\omega) = \lim_{T \rightarrow \infty} \frac{1}{T} \left\langle \iint_{\tau, v} x(v + \frac{\tau}{2}) x(v - \frac{\tau}{2}) e^{-j\omega \tau} d\tau dv \right\rangle$$

$$= \lim_{T \rightarrow \infty} \frac{1}{T} \iint_{\tau, v} C_x(\tau) e^{-j\omega \tau} d\tau dv$$



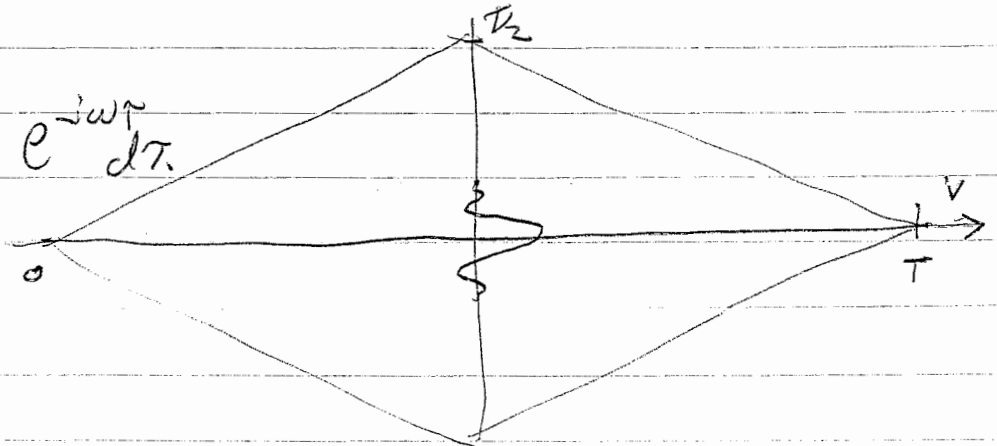
Once T is large enough

so that most of $C_x(\tau)$

has been covered, by, say $|\tau| \leq \frac{T}{1-\epsilon}$

then $\frac{1}{T} \int_0^T dv$ is within $\pm \epsilon$ of 1, where $C(\tau) \neq 0$.

$$\text{So } P_x(\omega) = \int_{-\infty}^{\infty} C_x(\tau) e^{-j\omega \tau} d\tau$$



13.

Analogous rels for bispectrum & cross-spectrum.

$$\left(\begin{array}{c} \downarrow \\ \langle x(\tau_1) x(\tau_2) x(-\tau_1 - \tau_2) \rangle \\ \downarrow \end{array} \right) \left(\begin{array}{c} \downarrow \\ \text{to cross-covariance} \\ \downarrow \end{array} \right)$$

Superposition property: If $x(t)$ & $y(t)$ are independent,

$$P_{x+y} = P_x + P_y$$

[cross-terms in det. of P.S. $\rightarrow 0$]

Weaker condition:

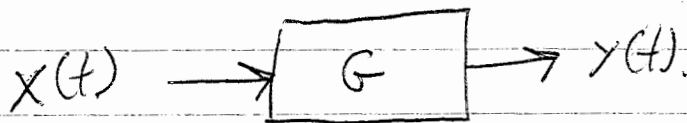
If $x(t)$ and $y(t')$ are uncorrelated - $P_{x+y} = P_x + P_y$.

Uncorrelated: $\langle x(t) y(t') \rangle = 0$

Independent: $\langle A(x) B(y) \rangle = \langle A(x) \rangle \langle B(y) \rangle$

In either circumstance, $C_{xy}(\omega) = 0$.

Rel. to linear filtering:



$$F(\omega, y) = \hat{G}(\omega) F(\omega, x)$$

$$P_y(\omega) = \lim_{T \rightarrow \infty} \frac{1}{T} \langle |F(y, \omega)|^2 \rangle = \lim_{T \rightarrow \infty} \frac{1}{T} \langle |\hat{G}(\omega) F(x, \omega)|^2 \rangle$$

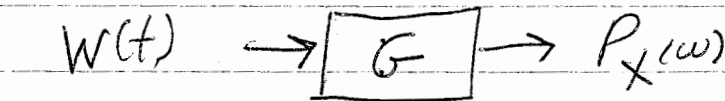
$$= |\hat{G}(\omega)|^2 P_x(\omega)$$

Implications:

$$|G(\omega)|^2 = \frac{P_y(\omega)}{P_x(\omega)}$$

Note phase of $G(\omega)$ cannot be determined

- Analysis of connectivity
- Can think of any Gaussian noise as a result of



$$\uparrow$$

$$P_W(\omega) = 1$$

provided $|G(\omega)| = \sqrt{P_x(\omega)}$

Though it may not be clear if phases can be assigned to $G(\omega)$ so that $\int G(\omega) e^{i\omega t} dt$ is causal & finite.